# CSC 311: Introduction to Machine Learning Lecture 7 - Probabilistic Models

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University of Toronto, Summer 2023





**2** Discriminative and Generative Classifiers

3 Naïve Bayes Models



- So far in the course we have adopted a modular perspective, in which the model, loss function, optimizer, and regularizer are specified separately.
- Today we begin putting together a probabilistic interpretation of our model and loss, and introduce the concept of maximum likelihood estimation.

#### 1 Probabilistic Modeling of Data

- 2 Discriminative and Generative Classifiers
- **3** Naïve Bayes Models
- 4 Bayesian Parameter Estimation

#### Example: A Biased Coin

You flip a coin N = 100 times and get outcomes  $\{x_1, \ldots, x_N\}$ where  $x_i \in \{0, 1\}$  and  $x_i = 1$  is interpreted as heads H.

Suppose you had  $N_H = 55$  heads and  $N_T = 45$  tails.

We want to create a model to predict the outcome of the next coin flip. That is, we want to answer this question:

What is the probability it will come up heads if we flip again?

#### Model

The coin is likely biased. Let's assume that one coin flip outcome x is a Bernoulli random variable for a currently unknown parameter  $\theta \in [0, 1]$ .

$$p(x = 1|\theta) = \theta$$
 and  $p(x = 0|\theta) = 1 - \theta$   
or more succinctly  $p(x|\theta) = \theta^x (1 - \theta)^{1-x}$ 

Assume that  $\{x_1, \ldots, x_N\}$  are independent and identically distributed (i.i.d.). Thus, the joint probability of the outcome  $\{x_1, \ldots, x_N\}$  is

$$p(x_1, ..., x_N | \theta) = \prod_{i=1}^N \theta^{x_i} (1 - \theta)^{1 - x_i}$$

The likelihood function is the probability of observing the data as a function of the parameters  $\theta$ :

$$L(\theta) = \prod_{i=1}^{N} \theta^{x_i} (1-\theta)^{1-x_i}$$

We usually work with log-likelihoods:

$$\ell(\theta) = \sum_{i=1}^{N} x_i \log \theta + (1 - x_i) \log(1 - \theta)$$

#### Maximum Likelihood Estimation

How can we choose  $\theta$ ? Good values of  $\theta$  should assign high probability to the observed data.

The maximum likelihood criterion says that we should pick the parameters that maximize the likelihood.

$$\hat{\theta}_{\mathrm{ML}} = \operatorname*{arg\,max}_{\theta \in [0,1]} \ell(\theta)$$

We can find the optimal solution by setting derivatives to zero.

$$\frac{\mathrm{d}\ell}{\mathrm{d}\theta} = \frac{\mathrm{d}}{\mathrm{d}\theta} \left( \sum_{i=1}^{N} x_i \log \theta + (1-x_i) \log(1-\theta) \right) = \frac{N_H}{\theta} - \frac{N_T}{1-\theta}$$

where  $N_H = \sum_i x_i$  and  $N_T = N - \sum_i x_i$ . Setting this to zero gives the maximum likelihood estimate:

$$\hat{\theta}_{\rm ML} = \frac{N_H}{N_H + N_T}$$

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- define a model that assigns a probability (or has a probability density at) to a dataset
- maximize the likelihood (or minimize the neg. log-likelihood).



#### 2 Discriminative and Generative Classifiers

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# Spam Classification

For a large company that runs an email service, one of the important predictive problems is the automated detection of spam email.





| Dear Karim,  |          |
|--|----------|
| I think we should postpone the board meeting to be held<br>after Thanksgiving.   | Not spam |
| Regards,<br>Anna   |          |
| Dear Toby,   |          |
| I have an incredible opportunity for mining 2 Bitcoin a day. Please<br>Contact me at the earliest at +1 123 321 1555. You won't want to miss<br>out on this opportunity. | Spam     |
| Regards,<br>Ark  |          |

# **Discriminative Classifiers**

**Discriminative** classifiers try to learn mappings directly from the space of inputs  $\mathcal{X}$  to class labels  $\{0, 1, 2, \dots, K\}$ 



# Generative Classifiers

**Generative** classifiers try to build a model of "what data for a class looks like", i.e. model  $p(\mathbf{x}, y)$ . If we know p(y) we can easily compute  $p(\mathbf{x}|y)$ .

Classification via Bayes rule (thus also called Bayes classifiers)



# Generative vs Discriminative

- Discriminative approach: estimate parameters of decision boundary/class separator directly from labeled examples.
  - Model  $p(t|\mathbf{x})$  directly (logistic regression models)
  - Learn mappings from inputs to classes (linear/logistic regression, decision trees etc)
  - ▶ Tries to solve: How do I separate the classes?
- Generative approach: model the distribution of inputs characteristic of the class (Bayes classifier).
  - Model  $p(\mathbf{x}|t)$
  - Apply Bayes Rule to derive  $p(t|\mathbf{x})$ .
  - ▶ Tries to solve: What does each class "look" like?
- Key difference: is there a distributional assumption over inputs?

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# Example: Spam Detection

- Classify email into spam (c = 1) or non-spam (c = 0).
- Binary features  $\mathbf{x} = [x_1, \dots, x_D], x_i \in \{0, 1\}$  saying whether each of D words appears in the e-mail.

Example email: "You are one of the very few who have been selected as a winner for the free \$1000 Gift Card."

Feature vector for this email:

```
..."card": 1
```

```
• ...
```

- "winners": 1
- "winter": 0
- ...
- "you": 1

# Bayesian Classifier

Given features  $\mathbf{x} = [x_1, x_2, \cdots, x_D]^T$ want to compute class probabilities using Bayes Rule:



In words,

 $Posterior for class = \frac{Pr. of feature given class \times Prior for class}{Pr. of feature}$ 

To compute  $p(c|\mathbf{x})$  we need:  $p(\mathbf{x}|c)$  and p(c).

# Motivation for Compact Representation

- Two classes:  $c \in \{0, 1\}$ .
- Binary features  $\mathbf{x} = [x_1, \dots, x_D], x_i \in \{0, 1\}$
- Define a joint distribution  $p(c, x_1, \ldots, x_D)$ . How many probabilities do we need to specify this joint dist.?
- Let's impose structure on the distribution so that the representation is compact and allows for efficient learning and inference

## Naïve Bayes Independence Assumption

Naïve assumption:

the features  $x_i$  are conditionally independent given the class c.

• Allows us to decompose the joint distribution:

$$p(c, x_1, \ldots, x_D) = p(c) p(x_1|c) \cdots p(x_D|c).$$

Compact representation of the joint distribution

- Prior probability of class:  $p(c = 1) = \pi$  (e.g. prob of spam)
- Conditional probability of feature given class:  $p(x_j = 1|c) = \theta_{jc}$  (e.g. prob of word appearing in spam)

#### Bayesian Network for a Naive Bayes Model



- Which probabilities do we need to specify this dist.?
- How many probabilities do we need to specify this dist.?

# Decomposing the Log-Likelihood

Decompose the log-likelihood into independent terms. Optimize each term independently.

$$\begin{split} \ell(\boldsymbol{\theta}) &= \sum_{i=1}^{N} \log p(c^{(i)}, \mathbf{x}^{(i)}) = \sum_{i=1}^{N} \log \left\{ p(\mathbf{x}^{(i)} | c^{(i)}) p(c^{(i)}) \right\} \\ &= \sum_{i=1}^{N} \log \left\{ p(c^{(i)}) \prod_{j=1}^{D} p(x_{j}^{(i)} | c^{(i)}) \right\} \\ &= \sum_{i=1}^{N} \left[ \log p(c^{(i)}) + \sum_{j=1}^{D} \log p(x_{j}^{(i)} | c^{(i)}) \right] \\ &= \sum_{\substack{i=1\\\text{Log-likelihood}\\\text{of labels}} \sum_{\substack{i=1\\\text{Log-likelihood}\\\text{for feature } x_{j}}} \log p(x_{j}^{(i)} | c^{(i)}) \end{split}$$

#### Learning the Prior over Class

- To learn the prior, we maximize  $\sum_{i=1}^{N} \log p(c^{(i)})$
- Define  $\pi = p(c^{(i)} = 1)$
- Pr. *i*-th email:  $p(c^{(i)}) = \pi^{c^{(i)}} (1-\pi)^{1-c^{(i)}}$ .
- Log-likelihood of the dataset:

$$\sum_{i=1}^{N} \log p(c^{(i)}) = \sum_{i=1}^{N} c^{(i)} \log \pi + \sum_{i=1}^{N} (1 - c^{(i)}) \log(1 - \pi)$$

• Maximum likelihood estimate of the prior  $\pi$  is the fraction of spams in dataset.

$$\hat{\pi} = \frac{\sum_{i} \mathbb{I}[c^{(i)} = 1]}{N} = \frac{\# \text{ spams in dataset}}{\text{total } \# \text{ samples}}$$

# Learning Pr. Feature Given Class

- To learn p(x<sub>j</sub><sup>(i)</sup> = 1 | c), we maximize Σ<sup>N</sup><sub>i=1</sub> log p(x<sub>j</sub><sup>(i)</sup> | c<sup>(i)</sup>)
  Define θ<sub>jc</sub> = p(x<sub>j</sub><sup>(i)</sup> = 1 | c).
- Pr. of *i*-th email:  $p(x_j^{(i)} | c) = \theta_{jc}^{x_j^{(i)}} (1 \theta_{jc})^{1 x_j^{(i)}}$ .
- Log-likelihood of the dataset:

$$\sum_{i=1}^{N} \log p(x_j^{(i)} | c^{(i)}) = \sum_{i=1}^{N} c^{(i)} \left\{ x_j^{(i)} \log \theta_{j1} + (1 - x_j^{(i)}) \log(1 - \theta_{j1}) \right\}$$
$$+ \sum_{i=1}^{N} (1 - c^{(i)}) \left\{ x_j^{(i)} \log \theta_{j0} + (1 - x_j^{(i)}) \log(1 - \theta_{j0}) \right\}$$

• Maximum likelihood estimate of  $\theta_{jc}$ is the fraction of word j occurrances in each class in the dataset.

$$\hat{\theta}_{jc} = \frac{\sum_{i} \mathbb{1}[x_{j}^{(i)} = 1 \& c^{(i)} = c]}{\sum_{i} \mathbb{1}[c^{(i)} = c]} \stackrel{\text{for } c = 1}{=} \frac{\# \text{word } j \text{ appears in class } c}{\# \text{ class } c \text{ in dataset}}$$

#### Predicting the Most Likely Class

- We predict the class by performing inference in the model.
- Apply Bayes' Rule:

$$p(c \mid \mathbf{x}) = \frac{p(c)p(\mathbf{x} \mid c)}{\sum_{c'} p(c')p(\mathbf{x} \mid c')} = \frac{p(c)\prod_{j=1}^{D} p(x_j \mid c)}{\sum_{c'} p(c')\prod_{j=1}^{D} p(x_j \mid c')}$$

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• For input **x**, predict c with the largest  $p(c) \prod_{j=1} p(x_j | c)$ (the most likely class).

$$p(c \mid \mathbf{x}) \propto p(c) \prod_{j=1}^{D} p(x_j \mid c)$$

- An amazingly cheap learning algorithm!
- Training time: estimate parameters using maximum likelihood
  - Compute co-occurrence counts of each feature with the labels.
  - Requires only one pass through the data!
- Test time: apply Bayes' Rule
  - Cheap because of the model structure. (For more general models, Bayesian inference can be very expensive and/or complicated.)
- Analysis easily extends to prob. distributions other than Bernoulli.
- Less accurate in practice compared to discriminative models due to its "naïve" independence assumption.

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Maximum likelihood can overfit if there is too little data.

Example: what if you flip the coin twice and get H both times?

$$\theta_{\rm ML} = \frac{N_H}{N_H + N_T} = \frac{2}{2+0} = 1$$

The model assigned probability 0 to T. This problem is known as data sparsity.

# Defining a Bayesian Model

We need to specify two distributions:

• The prior distribution  $p(\theta)$ encodes our beliefs about the parameters *before* we observe the data.

• The likelihood  $p(\mathcal{D} | \boldsymbol{\theta})$ encodes the likelihood of observing the data given the parameters. • When we update our beliefs based on the observations, we compute the posterior distribution using Bayes' Rule:

$$p(\boldsymbol{\theta} \mid \mathcal{D}) = \frac{p(\boldsymbol{\theta})p(\mathcal{D} \mid \boldsymbol{\theta})}{\int p(\boldsymbol{\theta}')p(\mathcal{D} \mid \boldsymbol{\theta}') \,\mathrm{d}\boldsymbol{\theta}'}.$$

- Rarely ever compute the denominator explicitly.
- In general, computing the denominator is intractable.

# Revisiting Coin Flip Example

We already know the likelihood:

$$L(\theta) = p(\mathcal{D}|\theta) = \theta^{N_H} (1-\theta)^{N_T}$$

It remains to specify the prior  $p(\theta)$ .

- An uninformative prior, which assumes as little as possible. A reasonable choice is the uniform prior.
- But, experience tells us 0.5 is more likely than 0.99. One particularly useful prior is the beta distribution:

$$p(\theta; a, b) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \, \theta^{a-1} (1-\theta)^{b-1}.$$

• We can ignore the normalization constant.

$$p(\theta; a, b) \propto \theta^{a-1} (1-\theta)^{b-1}.$$

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#### Beta Distribution Properties

- The expectation is  $\mathbb{E}[\theta] = a/(a+b).$
- The distribution gets more peaked when a and b are large.
- When a = b = 1, it becomes the uniform distribution.



## Posterior for the Coin Flip Example

• Computing the posterior distribution:

$$p(\boldsymbol{\theta} \mid \mathcal{D}) \propto p(\boldsymbol{\theta}) p(\mathcal{D} \mid \boldsymbol{\theta})$$
$$\propto \left[ \theta^{a-1} (1-\theta)^{b-1} \right] \left[ \theta^{N_H} (1-\theta)^{N_T} \right]$$
$$= \theta^{a-1+N_H} (1-\theta)^{b-1+N_T}.$$

A beta distribution with parameters  $N_H + a$  and  $N_T + b$ .

• The posterior expectation of  $\theta$  is:

$$\mathbb{E}[\theta \mid \mathcal{D}] = \frac{N_H + a}{N_H + N_T + a + b}$$

- Think of a and b as pseudo-counts. beta(a,b) = beta(1,1) + a - 1 heads + b - 1 tails.
- The prior and likelihood have the same functional form (conjugate priors).

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#### Bayesian Inference for the Coin Flip Example

When you have enough observations, the data overwhelm the prior.

Small data setting  $N_H = 2, N_T = 0$ 



Large data setting  $N_H = 55, N_T = 45$ 



## Maximum A-Posteriori (MAP) Estimation

Finds the most likely parameters under the posterior (i.e. the mode).



Converts the Bayesian parameter estimation problem into a maximization problem

$$\hat{\boldsymbol{\theta}}_{\text{MAP}} = \arg \max_{\boldsymbol{\theta}} p(\boldsymbol{\theta} \mid \mathcal{D})$$
$$= \arg \max_{\boldsymbol{\theta}} p(\boldsymbol{\theta}) p(\mathcal{D} \mid \boldsymbol{\theta})$$
$$= \arg \max_{\boldsymbol{\theta}} \log p(\boldsymbol{\theta}) + \log p(\mathcal{D} \mid \boldsymbol{\theta})$$

#### Maximum A-Posteriori Estimation

Joint probability of parameters and data:

$$\log p(\theta, \mathcal{D}) = \log p(\theta) + \log p(\mathcal{D} | \theta)$$
  
= Const + (N<sub>H</sub> + a - 1) log \theta + (N<sub>T</sub> + b - 1) log(1 - \theta)

Maximize by finding a critical point

$$\frac{\mathrm{d}}{\mathrm{d}\theta}\log p(\theta, \mathcal{D}) = \frac{N_H + a - 1}{\theta} - \frac{N_T + b - 1}{1 - \theta} = 0$$

Solving for  $\theta$ ,

$$\hat{\theta}_{\mathrm{MAP}} = \frac{N_H + a - 1}{N_H + N_T + a + b - 2}$$

## Estimate Comparison for Coin Flip Example

|                                 | Formula                                     | $N_H = 2, N_T = 0$       | $N_H = 55, N_T = 45$           |
|---------------------------------|---|--------------------------|--------------------------------|
| $\hat{	heta}_{\mathrm{ML}}$     | $\frac{N_H}{N_H + N_T}$                     | 1                        | $\frac{55}{100} = 0.55$        |
| $\mathbb{E}[	heta \mathcal{D}]$ | $\frac{N_H + a}{N_H + N_T + a + b}$         | $\frac{4}{6} pprox 0.67$ | $\frac{57}{104} \approx 0.548$ |
| $\hat{	heta}_{\mathrm{MAP}}$    | $\frac{N_H + a - 1}{N_H + N_T + a + b - 2}$ | $\frac{3}{4} = 0.75$     | $\frac{56}{102} \approx 0.549$ |

 $\hat{\theta}_{MAP}$  assigns nonzero probabilities as long as a, b > 1.